

Effect of Female Education and Labor Force Ratio on Economic Development

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Abstract

This study investigates the effects of women's education and labor force ratio on the level of development in countries. We use a complete dataset covering 44 countries over the period 1990–2010. It comprises the following: education index, the ratio of girls to boys in primary and secondary education, income per capita, human development index, life expectancy at birth, population in the 0–14 and 15–64 age intervals, population over 64, number of seats held by women in parliaments, and female labor force ratio. We interpret this dataset using panel regression analysis, with the human development index as the depend ent variable. The variables are also subject to factor analysis and this results in three statistically significant factors—education and development, female labor force ratio, and population. We examine the rankings of factor scores for the countries covered and relevant to the issues in question. We then interpret their consequences. The countries are ranked using factor analysis and the results of the fixed effects model, which studies the changes that occurred between 1990 and 2010.

Keywords: Female education ratio • Female labor force ratio • Human development index • Factor analysis • Fixed effects model

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One of the most significant problems in developing countries is the level of education. Rapid population growth in such countries has declined the schooling rate and quality of education, besides leading to shortages in qualified labor.

Rapid population growth has also led to the problem of integrating a qualified female population into the production process of these countries. Unlike developed countries, women in less developed countries are constrained by the notions surrounding the concept of a traditional family. Hence, despite being educated, women face barriers to entering the labor force and participating in the production process.

This study seeks the "woman" factor in human development and investigates if this factor has any impact on economic development. Hence, only countries with complete records for the period under study are included. The data are subject to panel regression and factor analyses. After interpreting the panel regression model defined by such variables as the education ratios of girl students, female labor force ratio, and number of seats held by women in parliament, we factorize the multiple variables through dimension reduction and factor analyses. They reveal the factor scores as well as the rankings and groupings of countries. In addition, we study country rankings emerging from both analyses and from changes in the countries examined over the period under consideration. However, first, we will briefly examine the concept of development.

Economic Development

Economic development refers to improvements in cultural and political spheres besides growth of the economy. It can also be defined as the enhancement of welfare. In other words, the term development embraces improvements in socio-cultural spheres and in freedoms along with improvement in statistical indicators. For example, while the growth of a child may refer to his learning, his ability to read and write, his transitioning into adolescence, and to physical changes such as the growth of facial hair, it is his intelligence and contribution to his environment that is referred to as his development.

Although a concept of development that encompasses all countries is impossible, certain core factors can be isolated to illustrate the meaning of development (Case, Fair, & Oster 2012, p. 715).

Some of the resources of economic development are capital accumulation, human resources,

entrepreneurship, social capital, education, and the female labor force. In this study, we will consider only education and the female labor force.

Education: One of the most crucial factors in economic development is education. Education is the process of effecting a desired change in the behaviors of an individual through personal experience. Developing countries use educational policies in developmental strategies (Seyidoğlu, 1993, p. 27). There is a strong correlation between development and education. Thus, countries that value the education of girls will rank higher in terms of development.

Female Labor Force: Women are a significant factor in the dynamics and regulation of societies. Besides their role in shaping social life, women also play an important role in driving the development and advancement of countries. Worldwide, statistics show that there is a positive correlation between the value placed on women and the development of countries. Development is seen in those countries where women are appreciated and are active participants in the economic, social, and political spheres. In contrast, one sees poverty and misery in countries that do not view women favorably.

The contribution of women to the labor force dates back to the Industrial Revolution (1760–1840). As the demand for workers increased and as the male labor force fell short of demand, the need for women workers increased. An increase in the education of women and their participation in the labor force ensured rapid development of countries. Hence, it is important to take measures to improve the education of women and increase their contribution to the labor force (Özdemir, Yalman, & Bayraktar, 2012). However, there is a significant difference between women's participation in employment as individuals with high-end positions or with worker status when one considers such positions as the indicators of the development of a country.

Statistical Methods

Panel Data Analysis

In econometric analyses, data can be divided into three groups: time series, cross-section, and mixed. If the variables are observed within a certain time period for a certain observation group, then such mixed data are called panel data. In other words, panel data combines sectional and time series analyses. Due to technology, panel data analysis has become a widely used mode of analysis. In such analysis, if there is no missing observation in the dataset, which means the dataset dimensions are the same for all countries, the dataset is called a *balanced panel*. However, if there are missing observations, and the dataset dimensions are not the same for all cases of observation, the dataset is called an *unbalanced panel*. The panel data model is given in equation 1 as follows:

$$\begin{split} Y_{it} &= \beta_{1it} + \beta_{2it} X_{2it} + \dots + \beta_{kit} X_{kit} + \Theta_{it} \\ &= 1, \dots, N \\ t &= 1, \dots, T \end{split}$$
 (1)

In this equation, N denotes units and t denotes time. When variable Y is the dependent variable, with different values from unit to unit and from one time period to another, two sub-indices—i and t—are used to denote the section (observation) dimension and time period, respectively. It is possible to obtain different models by making different assumptions, particularly with respect to the properties of error terms and variability of coefficients, in studies conducted through panel data. Models obtained by making different assumptions are called models with "fixed effects" and "random effects." It is assumed that e_n errors are distributed as N (0, σ 2), independent in both models for all time periods and individuals

While working with a panel dataset, various tests are administered to decide between fixed effects and random effects. The test statistic proposed by Hausman enables one to study the validity of the notion that error term components of a random effects model are not linked with the independent variables in the model. The Hausman model specification test helps one decide between fixed effects and random effects model for panel data analyses (Greene, 2003, pp. 72-73). It is important to study whether the difference between parameter estimators of a fixed effects model and those of a random effects model is statistically significant. While one prefers the high values of the Hausman statistic in a fixed effects model, lower values are preferred in random effects models. The Hausman test statistic tests the correlation between individual effects unique to cross-sectional (ɛi) and explanatory variables. This test is asymptotically $\chi 2$ distributed. Rejection of the null hypothesis makes one conclude that it is necessary to accept the fixed effects model vis-à-vis the random effects model. The following hypotheses can be defined:

H0: $E(\epsilon i | Xit) = 0$ Effects regarding the country and time are random.

There is no correlation between (ϵi) and explanatory variables.

H1: $E(\epsilon i | Xit) \neq 0$ Country and time effects are fixed.

There is a correlation between (ɛi) and explanatory variables.

The null hypothesis shows there is no correlation between the explanatory variables in the model and error terms. It is more appropriate to use random effects model in cases where the null hypothesis is accepted. In contrast, it is appropriate to use a fixed effects model if the null hypothesis is rejected. The Hausman test statistic implies a chi square distribution with k degrees of freedom under the null (0) hypothesis, "random effects estimator is acceptable." Upon its realization, one can decide that the error term components of a random effects model are uncorrelated with the independent variables. In such as case, a fixed effects model is chosen (Turhan & Taşseven, 2010, p. 142).

Factor Analysis

Factor analysis is one of the most widely used multivariate statistical techniques that transforms interrelated multiple variables into fewer and more comprehensible significant factors independent of one another.

Factor analysis helps one identify representative variables to be used in other analyses, among many other variables. Factors obtained from the original dataset are used in a wide range of methods, including regression, correlation, and discriminant analysis as well as interdependent methods such as cluster analysis (Hair, Anderson, Tatham, & Black, 1998, pp. 95-97). Factor analysis serves the function of revealing unobservable and immeasurable latent dimensions that lie behind many measurable and observable attributes.

Making classical assumptions such as normality, multicollinearity, and linearity leads to decline in correlations between variables in factor analysis. The normality assumption is required only if the significance of the factors to be derived will be tested. The assumption of multiple linear relations is desired as factor analysis reckons with the interrelationship of variables. Factor analysis is also appropriate if the correlation among variables is greater than 0.30.

The more the correlation among variables is, the more likely it is for the variables to form common factors (Kalaycı, 2006, pp. 321-322). Hence, besides being a basic statistical principle, factor analysis is also indicated when the correlation matrix has adequate significant correlations. It is acceptable to

extract variables, with correlation coefficients less than 30%, using factor analysis.

The Kaiser–Mayer–Olkin (KMO) test is a sampling adequacy test used for measuring the correlations among variables and the applicability of factor analysis. The test value ranges between 0 and 1. The KMO value equals 1 if any variable is estimated by other variables without any error. As shown below, the KMO test is done by comparing the calculated simple correlation coefficients with partial correlation coefficients.

$$KMO = \frac{\sum_{i \neq j} \sum r_{j}^{2}}{\sum_{i \neq j} \sum r_{j}^{2} + \sum_{i \neq j} \sum a_{j}^{2}}$$
(1-1)

where KMO denotes the Kaiser–Mayer–Olkin sampling adequacy test, r_{ij} denotes the simple correlation coefficient between the variables i and j, and a_{ij} denotes the partial correlation coefficient between the variables i and j. When the relation between the variable pairs cannot be explained by other variables, the KMO value decreases. In such a case, factor analysis should not be used. By providing a measure of the average correlation between variables, the test measures the homogeneity of the variables. For an outstanding factor analysis, the KMO value should not be greater than 0.80. However, values greater than 0.50 are acceptable.

Factor analysis uses the standardized Z_{pxn} data matrix, which is obtained from the X_{pxn} raw data matrix that gives the p number of variables of n individuals. In this case, the factor analysis model is a linear model that denotes the relation between z_j variables and $f_1, f_2, ..., f_m$ common factors (Tatlıdil, 2002, p. 168) and gives the correlation with the highest value (Pazarlıoğlu, Emeç, & Erdoğan, 1999, p. 850). This model is generally depicted as follows:

$$z_{j} = a_{j1}f_{1} + a_{j2}f_{2} + \ldots + a_{jm}f_{m} + b_{j}u_{j}$$
(1-2)

where $z_i : j^{th}$ variable,

- f: common factor
- u; specific or residual factor
- b: coefficient related to specific or residual factor
- m = the number of common factors

The value of j variable for i individual can be written as given below, following the classical factor

analysis model:

$$z_{j} = \sum_{p=1}^{m} a_{j} F_{j} + b_{j} u_{j} (i = 1, 2, ..., N; j = 1, 2, ..., n) (1-3)$$

This model can be clearly written as follows for any z_j variable depending on m common factors and a unique factor:

$$z_n = a_{n1}F_1 + a_{n2}F_2 + \dots + a_m F_m + b_n U$$

This set of equations is called a *factorial pattern* or just *pattern*. Here, the common factors F_p (p = 1, 2,..., m) may or may not be correlated. Yet, the unique factor U_j (j = 1,2,..., n) is always accepted as uncorrelated both with themselves and with other common factors. Factor analysis does not only yield a pattern but also produces the relationship between variables and factors. The table depicting this relation is called a *factorial structure* or just *structure*, and is required for completing both the pattern and the structural resolution (Atan, Göksel, & Karpat, 2002)

The factors are denoted and interpreted by considering factor loadings. Thus, if the correlation of the variable with the common factor is greater than ± 0.30 , it is accepted that the variable has significant correlation with the common factor in question. Some researchers say this proportion should be greater than ± 0.40 or even ± 0.50 . Negative factor weights should also be while including positive factor loadings that are greater than ± 0.30 . The common factor should be interpreted as showing a positive relation with positive factor weights and a negative relation with negative factor weights. Hence, the factors can be viewed as being dipolar (Albayrak, 2000, pp. 124-125).

Analysis and Findings

We looked at panel data for 44 countries over the period 1990 and 2010 to examine the effects of women's education and labor force ratio on the development of these countries. The data were obtained annually from the databanks of the United Nations Development Program (UNDP) and the World Bank. An optimal model was chosen using the E-Views 07 econometric analysis software. In the optimal model, which can be accessed in full for 12 variables, three were found to be statistically significant (p < .001). Meanwhile, factor analysis

was also applied to the 12 variables to see the profile of all the variables. Thus, factor scores helped to compare the ranking of countries, in terms of the fixed effect, with the sequence. Significant factorizations resulted from applying factor analysis to the 12 identified variables. The countries were ranked on the basis of the following factors:

The variables used in the study are defined as follows:

- X1: Education index
- X2: The ratio of girls to boys in primary education
- X3: The ratio of girls to boys in secondary education
- X4: Income per capita (in dollars)
- X5: Human development index*
- X6: Life expectancy at birth

X7: The ratio of population aged 0–14 to total population

X8: The ratio of population aged 15–64 to total population

X9: Population aged over 64

X10: Total population

X11: Proportion of seats held by women in national parliaments (%)

X12: Ratio of female to male labor participation rate (%)

Factor analyses were applied separately to 12 variables that could explain development for five different periods (1990, 1995, 2000, 2005, and 2010) to explain development. A factorization outcome was obtained for each period. The results revealed three factors with which the variables could coexist. These factors, given in Table 1, were education and development, female labor force ratio, and the general population. This classification was based on a study of the factor components. As similar loadings were found for the five different periods studied, only the component matrix obtained from the data for the year 2010 is provided in Table 1.

Rankings based on factor loadings in the countries studied are given in Table 2. As the variables compiled in the first factor are linked to education and development, the ranking of countries are based on such factor loadings of the countries. Countries are ranked based on separate factor loadings for each year. It is possible to compare changes in countries over the years in one single table where such rankings are indicated. The European Union member countries are indicated in bold. The table also shows the proximity of other countries to EU member countries.

Table 1 Component Matrix of 2010				
		Component		
	1	2	3	
x1	.966	.065	.007	
x7	950	.048	.261	
x6	.929	225	081	
x4	.888	.283	090	
x9	.872	.242	217	
x8	.833	356	253	
x12	049	.902	.030	
x11	.319	.691	.271	
x10	.121	319	.304	
x3	.483	248	.698	
x2	.487	.018	.696	

We used E-Views 07 software to compile data for the panel regression analysis. The significant model, with variables related to female rates as independent variables and with the ones related to human development index as dependent variables, is given in Table 3 and interpreted.

Analyzing the significance of coefficients in the model equation, we can draw the following inferences: The contribution of the variable X2, which is the ratio of girls to boys in primary education, to the model equation is insignificant. It is included in the model because of higher probability of type II error of this variable. While the ratio of girls to boys in secondary education (X3) is significant, the ratio of girls to boys in primary education (X2) is also intended to be shown as approximately significant. The contribution of other variables to the model equation is significant.

The United Nations Development Program has been publishing an annual report called the Human development report since1990. The UNDP's Human Development Index helps compare social and economic development levels across countries. In its report, the UNDP states that it is inappropriate to measure the level of development across countries based on an increase in income per capita (economic growth). A high national income is not enough to call a country developed. Social and economic developments are better measures of welfare and living standards across countries than economic growth.

1990	1995	2000
Norway	Japan	Sweden
Sweden	Sweden	Norway
witzerland	United States	Belgium
Inited States	Norway	Spain
Finland	Belgium	Japan
letherlands	Italy	Finland
Japan	Switzerland	United States
Belgium	Spain	Netherlands
Italy	Finland	Switzerland

1990	1995	2000	2005	2010
Norway	Japan	Sweden	Norway	Norway
Sweden	Sweden	Norway	Spain	Japan
Switzerland	United States	Belgium	Sweden	Sweden
United States	Norway	Spain	Japan	Switzerland
Finland	Belgium	Japan	Finland	Netherlands
Netherlands	Italy	Finland	United States	Spain
Japan	Switzerland	United States	Switzerland	Finland
Belgium	Spain	Netherlands	Australia	United States
Italy	Finland	Switzerland	Austria	Austria
Australia	Netherlands	Australia	Belgium	Australia
Austria	Australia	Austria	Italy	Belgium
France	Austria	Italy	Netherlands	Italy
Spain	France	New Zealand	New Zealand	New Zealand
New Zealand	Greece	Greece	Greece	Korea (Republic of)
Greece	New Zealand	France	France	France
Hungary	Portugal	Portugal	Portugal	Greece
Bulgaria	Hungary	Malta	Korea (Republic of)	Portugal
Portugal	Korea (Republic of)	Korea (Republic of)	Cyprus	Cyprus
Cyprus	Cyprus	Hungary	Malta	Malta
Malta	Bulgaria	Cyprus	Hungary	Hungary
Korea (Republic of)	Malta	Bulgaria	Bulgaria	Bulgaria
Uruguay	Uruguay	Uruguay	Costa Rica	Costa Rica
Colombia	Costa Rica	Costa Rica	Uruguay	Uruguay
Costa Rica	Dominica	Mexico	Tunisia	Tunisia
Mexico	Mexico	Tunisia	Mexico	Mexico
Dominica	Colombia	Dominica	Algeria	Algeria
Paraguay	Tunisia	Colombia	Dominica	Colombia
Indonesia	Indonesia	Indonesia	Colombia	Turkey
Nicaragua	Paraguay	Algeria	Indonesia	Indonesia
Honduras	Turkey	Paraguay	Turkey	Dominica
Tunisia	Nicaragua	Turkey	Paraguay	Paraguay
Turkey	Honduras	Nicaragua	Honduras	Egypt
Egypt	Algeria	Honduras	Egypt	Syrian Arab Rep.
Algeria	Egypt	Egypt	Nicaragua	Honduras
Syrian Arab Rep.	Syrian Arab Rep.	Syrian Arab Rep.	Syrian Arab Rep.	Nicaragua
Madagascar	Madagascar	Morocco	Morocco	Morocco
Morocco	Morocco	Madagascar	Madagascar	Senegal
Lao People's D.Rep.	Lao People's D.Rep.	Lao People's D.Rep.	Senegal	Madagascar
Uganda	Malawi	Senegal	Lao People's D.Rep.	Lao People's D.Rep.
Senegal	Senegal	Uganda	Uganda	Uganda
Malawi	Uganda	Malawi	Malawi	Malawi
Togo	Togo	Togo	Togo	Togo
Mozambique	Mozambique	Mozambique	Guinea	Mozambique
Guinea	Guinea	Guinea	Mozambique	Guinea

Parameters are estimated by using both the fixed effects and random effects model to understand the effects of observation in panel data. The Hausman test is performed to decide which of the two models is statistically valid. In the Hausman test, the null hypothesis is that the preferred model is the "random effects model," whereas the alternative hypothesis is that the preferred model is the "fixed effects model." Table 4 gives the Hausman test results.

Table 3 Outcome of Panel Regression	ı Analysis			
Dependent Variable: X5				
Method: Pooled Least Square	res			
Sample: 1990-2010				
Included observations: 4				
Cross-sections included: 44				
Total pool (balanced) obser	vations: 176			
Variable	Coefficient	Std. Error	t-Statistic	Prob.
X2	0.002278	0.001555	1.464.839	0.1448
X3	0.005413	0.001470	3.683.261	0.0003
X11	0.005275	0.001103	4.783.640	0.0000
X12	-0.002035	0.000533	-3.815.853	0.0002
R-squared	0.476646	Mean dependent	var	0.686199
Adjusted R-squared	0.467518	S.D. dependent v	ar	0.188961
S.E. of regression	0.137887	Akaike info criter	rion	-1.102.298
Sum squared residuals	3.270.212	Schwarz criterion	L	-1.030.241
Log likelihood	1.010.022	Hannan–Quinn d	criterion	-1.073.072
Durbin-Watson statistic	0.104900			

X5 = 0.002278 X2 + 0.005413X3 + 0.005275X11 - 0.002035X12

Table 4 Hausman Test Results				
Correlated	Random Ef	fects - Hausm	an Test	
Pool: HDI				
Test cross-	section rand	om effects		
Test Summary Chi-Sq. Chi-Sq. d.f. Prob.				
Cross-section ran- dom		48.148.782	4	0.0000
Cross-section random effects test comparisons:				
Variable Fixed		Random	Var (Diff.)	Prob.
X2	0.003363	0.003424	0.000000	0.7693
X3	0.001440	0.001692	0.000000	0.0865
X11	0.001681	0.002872	0.000000	0.0000
X12	0.003713	0.001499	0.000000	0.0000

As probability = .000 < .05 is seen, H₀ hypothesis is rejected, which means the model will be estimated through the fixed effects.

Table 5 gives the fixed effects estimation results.

Analyzing the coefficients of the model, we find that each of the variables makes a significant contribution to the fixed effect model. Table 6 gives the ranking of countries, from low to high, based on the model's fixed effects value.

Discussion

Economic development, or the enhancement of welfare, refers to improvements in the social, cultural, and political spheres along with economic growth. This study reveals the effect of women's education and labor force participation on the economic and human development levels of countries statistically over a period of roughly 20 years, or 1990 to 2010.We examined 44 countries, whose data could be accessed for the given time period. The 12 variables of the component scores were compiled into three factors based on the results of factor analysis. This result, yielded as the same factorization for five different years, demonstrating that the correlation among the developmental indicators did not change over the years. The following variables were compiled for the first factor, seen as the most important: education index (x1), the ratio of population in the 0-14 age interval to total population (x7), life expectancy at birth (x6), income per capita (x4), population aged over 64 (x9), and the ratio of population in the 15–64 age interval to total population (x8). Among these, only the ratio of population in the 0-14 age interval to total population had a negative impact (-0.95) on the relevant factor. The loadings of other variables were positive with rather high ratios (see Table 3). Numerical values show that an increase in the population of children is inversely correlated with the indicators of development. As for the rankings of countries based on the scores of this factor, no marked change was observed over the years (see Table 4). However, Tunisia, Korea, and Spain displayed marked differences. All these three countries made seven levels of progress between 1990 and 2010. Similarly, Algeria made a progress of eight levels. Colombia, Nicaragua, Honduras, and Dominica were the countries that regressed in the rankings between 1990 and 2010. Turkey made a progress of four levels. Overall, members of the

Table 5						
	ects Estimation Results					
	nt Variable: X5					
-	Pooled Least Squares					
Sample: 1	-					
	observations: 4					
	tions included: 44					
	l (balanced) observations	. 176				
Variable	Coefficient	Std. Error	t-Statistic	Prob.	-	
C	-0.552255	0.136442	-4.047.557	0.0001		
X2						
	0.008161	0.002083	3.917.556	0.0001		
X3	0.004864	0.001436	3.387.571	0.0009		
X11	0.004236	0.001153	3.672.605	0.0003		
X12	-0.001205	0.000554	-2.176.637	0.0309		-
Fixed Effe	ects (Cross)					-
	_ALGERIAC	-0.0	050708			
	_AUSTRALIAC	0.2	23283		Effects Specifica	ation
	_AUSTRIAC	0.1	68468		-	
	_ _BELGIUMC		84855			
	_BULGARIAC		51394	Cross-section fixed	(dummy variable	es)
	COLOMBIAC		043155	Period fixed (dumn	•	
1	_COSTARICAC		29142			
	_CYPRUSC		10632			
	_DOMINICAC		02666	R-squared	0.989069	Mean dependent var
	_EGYPTC		095694	Adjusted R-squared		S.D. dependent var
	_FINLANDC		76038	S.E. of regression	0.023375	Akaike info criterion
	_FRANCEC		.53544	Sum squared re-		Schwarz criterion
	_ _GREECEC		28102	sidual Log likelihood	4.414.464	Hannan–Quinn criterion
	_GUINEAC		355530	F-statistic	2.262.139	Durbin-Watson statistic
	_HONDURASC			Prob (F-statistic)	0.000000	Durbhi – watson statistic
			124159	PIOD (F-statistic)	0.000000	
	_HUNGARYC INDONESIAC		191094			
	-		139496			
	_ITALYC		42734	Fixed Effects	(Dariad)	1
V	_JAPANC		.75607			
_	OREAREPUBLICC		41052	1990C	066161	
_	AOPEOPLED.REPC		193761	2000C	003355	
	MADAGASCARC		268260	2005C	0.023023	
	_MALAWIC		324793	2010C	0.046493]
	_MALTAC		07122			
	_MEXICOC		31487			
	_MOROCCOC		150376			
	MOZAMBIQUEC		373395			
1	NETHERLANDSC		12725			
	NEWZEALANDC		02009			
-	_NICARAGUAC	-0.	153831			
	_NORWAYC	0.2	43860			
	_PARAGUAYC	-0.0	074418			
	_PORTUGALC	0.0	88239			
	_SENEGALC	-0.2	255158			
	_SPAINC	0.1	54240			
	_SWEDENC	0.2	13341			
	SWITZERLANDC	0.1	97393			
_	YRIANARABREPC		084606			
	_TOGOC		238756			
	_TUNISIAC		039439			
	TURKEYC		034268			
	_UGANDAC		290358			
τ	JNITEDSTATESC		16330			
_`	_URUGUAYC		44805			
		0.0		L		

Table The I	e 6 Ranking of Countries	
1	_MOZAMBIQUEC	-0.3734
2	_GUINEAC	-0.35553
3	_MALAWIC	-0.32479
4	_UGANDAC	-0.29036
5	_MADAGASCARC	-0.26826
6	_SENEGALC	-0.25516
7	_TOGOC	-0.23876
8	_LAOPEOPLEDEMREPC	-0.19376
9	_NICARAGUAC	-0.15383
10	_MOROCCOC	-0.15038
11	_INDONESIAC	-0.1395
12	_HONDURASC	-0.12416
13	_EGYPTC	-0.09569
14	_SYRIANARABREPUBLIC	-0.08461
15	_PARAGUAYC	-0.07442
16	_ALGERIAC	-0.05071
17	_COLOMBIAC	-0.04316
18	_TUNISIAC	-0.03944
19	_TURKEYC	-0.03427
20	_DOMINICAC	0.002666
21	_UNITEDSTATESC	0.01633
22	_COSTARICAC	0.029142
23	_MEXICOC	0.031487
24	_URUGUAYC	0.044805
25	_BULGARIAC	0.051394
26	_PORTUGALC	0.088239
27	_HUNGARYC	0.091094
28	_MALTAC	0.107122
29	_CYPRUSC	0.110632
30	_GREECEC	0.128102
31	_KOREAREPUBLICC	0.141052
32	_ITALYC	0.142734
33	_FRANCEC	0.153544
34	_SPAINC	0.15424
35	_AUSTRIAC	0.168468
36	_JAPANC	0.175607
37	_FINLANDC	0.176038
38	_BELGIUMC	0.184855
39	_SWITZERLANDC	0.197393
40	_NEWZEALANDC	0.202009
41	_NETHERLANDSC	0.212725
42	_SWEDENC	0.213341
43	_AUSTRALIAC	0.223283
44	_NORWAYC	0.24386

European Union, the USA, Japan, and Australia enjoyed high rankings, which was a notable finding.

The results of the panel regression analysis show the insignificance of the ratio of girls in primary education in the model. This can be interpreted as primary education, compulsory in most countries, losing its discriminative effect on development. Although this variable was found to be insignificant, the ratio of girls to boys in secondary education turned out to be a significant variable in the model. Thus, if girls continue their education, it will have an effect on development. The negative value (β = -.002; p < .001) of the coefficient related to the ratio of women workers was found to be significant in the model, indicating that an increase in the rate of participation of women laborers is inversely correlated with human development. Thus, women are employed as laborers in underdeveloped countries. However, the more developed a country is, the more are the women employed in higher-end jobs. There are great similarities in the ranking of countries based on the results of factor analysis and according to the fixed-effect coefficients, even when the variables vary. Developed countries occupy the top row, while the less developed countries populate the bottom row. As for Turkey, she keeps the same rank after each analysis. Therefore, irrespective of the variables supporting women's education, it is still an indicator of the level of development of a country.

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